

## Section 7.1

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# Conditioning a Random Variable by an Event

## Example 7.1

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Let  $N$  equal the number of bytes in an email. A conditioning event might be the event  $I$  that the email contains an image. A second kind of conditioning would be the event  $\{N > 100,000\}$ , which tells us that the email required more than 100,000 bytes. Both events  $I$  and  $\{N > 100,000\}$  give us information that the email is likely to have many bytes.

## **Definition 7.1 Conditional CDF**

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*Given the event  $B$  with  $P[B] > 0$ , the conditional cumulative distribution function of  $X$  is*

$$F_{X|B}(x) = P[X \leq x|B].$$

# Conditional PMF Given an

## **Definition 7.2** Event

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*Given the event  $B$  with  $P[B] > 0$ , the conditional probability mass function of  $X$  is*

$$P_{X|B}(x) = P[X = x|B].$$

# Conditional PDF Given an

## **Definition 7.3** Event

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*For a random variable  $X$  and an event  $B$  with  $P[B] > 0$ , the conditional PDF of  $X$  given  $B$  is*

$$f_{X|B}(x) = \frac{dF_{X|B}(x)}{dx}.$$

# Theorem 7.1

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For a random variable  $X$  and an event  $B \subset S_X$  with  $P[B] > 0$ , the conditional PDF of  $X$  given  $B$  is

$$\text{Discrete: } P_{X|B}(x) = \begin{cases} \frac{P_X(x)}{P[B]} & x \in B, \\ 0 & \text{otherwise} \end{cases}$$

$$\text{Continuous: } f_{X|B}(x) = \begin{cases} \frac{f_X(x)}{P[B]} & x \in B, \\ 0 & \text{otherwise.} \end{cases}$$

## Example 7.2 Problem

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A website distributes instructional videos on bicycle repair. The length of a video in minutes  $X$  has PMF

$$P_X(x) = \begin{cases} 0.15 & x = 1, 2, 3, 4, \\ 0.1 & x = 5, 6, 7, 8, \\ 0 & \text{otherwise.} \end{cases} \quad (7.3)$$

Suppose the website has two servers, one for videos shorter than five minutes and the other for videos of five or more minutes. What is the PMF of video length in the second server?

## Example 7.2 Solution

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We seek a conditional PMF for the condition  $x \in L = \{5, 6, 7, 8\}$ . From Theorem 7.1,

$$P_{X|L}(x) = \begin{cases} \frac{P_X(x)}{P[L]} & x = 5, 6, 7, 8, \\ 0 & \text{otherwise.} \end{cases} \quad (7.4)$$

From the definition of  $L$ , we have

$$P[L] = \sum_{x=5}^8 P_X(x) = 0.4. \quad (7.5)$$

With  $P_X(x) = 0.1$  for  $x \in L$ ,

$$P_{X|L}(x) = \begin{cases} 0.1/0.4 = 0.25 & x = 5, 6, 7, 8, \\ 0 & \text{otherwise.} \end{cases} \quad (7.6)$$

Thus the lengths of long videos are equally likely. Among the long videos, each length has probability 0.25.



## Example 7.3 Problem

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For the pointer-spinning experiment of Example 4.1, find the conditional PDF of the pointer position for spins in which the pointer stops on the left side of the circle.

## Example 7.3 Solution

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Let  $L$  denote the left side of the circle. In terms of the stopping position,  $L = [1/2, 1)$ . Recalling from Example 4.4 that the pointer position  $X$  has a uniform PDF over  $[0, 1)$ ,

$$P[L] = \int_{1/2}^1 f_X(x) dx = \int_{1/2}^1 dx = 1/2. \quad (7.7)$$

Therefore,

$$f_{X|L}(x) = \begin{cases} 2 & 1/2 \leq x < 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.8)$$

## Example 7.4 Problem

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Suppose  $X$ , the time in integer minutes you wait for a bus, has the discrete uniform PMF

$$P_X(x) = \begin{cases} 1/20 & x = 1, 2, \dots, 20, \\ 0 & \text{otherwise.} \end{cases} \quad (7.9)$$

Suppose the bus has not arrived by the eighth minute; what is the conditional PMF of your waiting time  $X$ ?

## Example 7.4 Solution

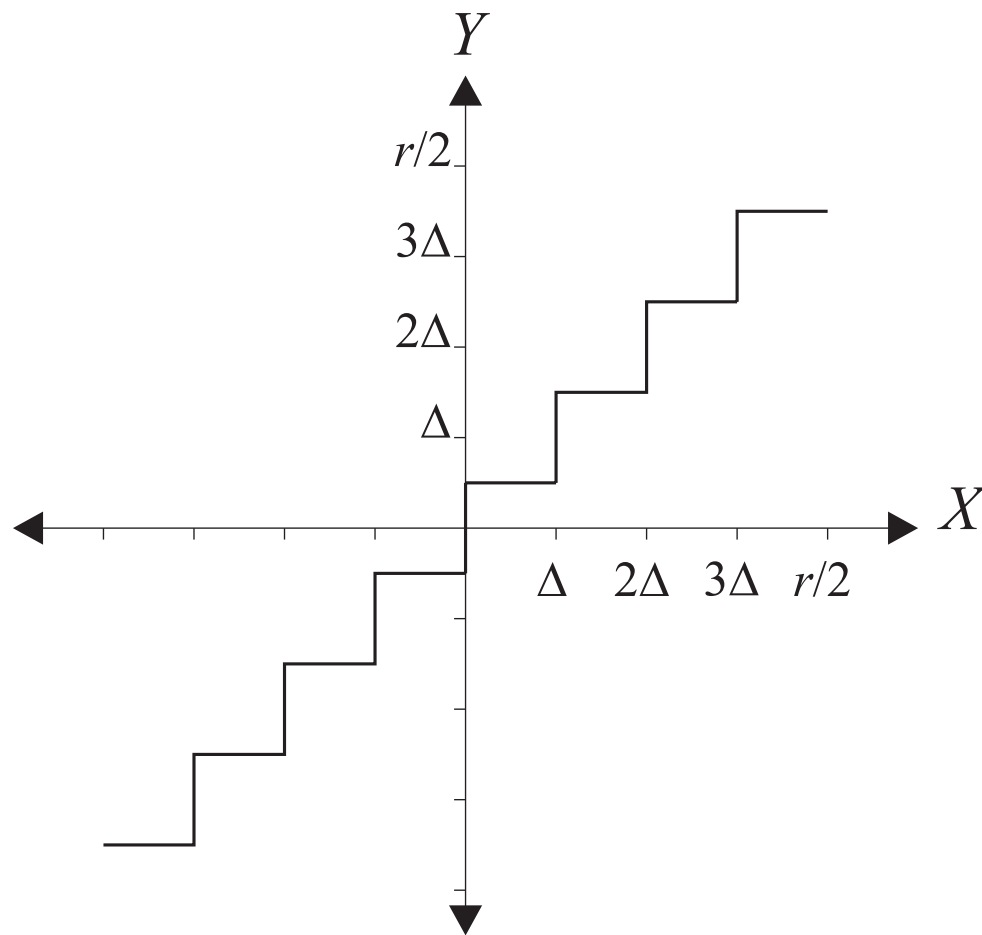
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Let  $A$  denote the event  $X > 8$ . Observing that  $P[A] = 12/20$ , we can write the conditional PMF of  $X$  as

$$P_{X|X>8}(x) = \begin{cases} \frac{1/20}{12/20} = \frac{1}{12} & x = 9, 10, \dots, 20, \\ 0 & \text{otherwise.} \end{cases} \quad (7.10)$$

## Example 7.5 Problem

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The continuous uniform  $(-r/2, r/2)$  random variable  $X$  is processed by a  $b$ -bit uniform quantizer to produce the quantized output  $Y$ . Random variable  $X$  is rounded to the nearest quantizer level. With a  $b$ -bit quantizer, there are  $n = 2^b$  quantization levels. The quantization step size is  $\Delta = r/n$ , and  $Y$  takes on values in the set

$$Q_Y = \{y_{-n/2}, y_{-n/2+1}, \dots, y_{n/2-1}\} \quad (7.11)$$

where  $y_i = \Delta/2 + i\Delta$ . This relationship is shown for  $b = 3$  in the figure on the left. Given the event  $B_i$  that  $Y = y_i$ , find the conditional PDF of  $X$  given  $B_i$ .

## Example 7.5 Solution

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In terms of  $X$ , we observe that  $B_i = \{i\Delta \leq X < (i+1)\Delta\}$ . Thus,

$$\mathbb{P}[B_i] = \int_{i\Delta}^{(i+1)\Delta} f_X(x) dx = \frac{\Delta}{r} = \frac{1}{n}. \quad (7.12)$$

By Definition 7.3,

$$f_{X|B_i}(x) = \begin{cases} \frac{f_X(x)}{\mathbb{P}[B_i]} & x \in B_i, \\ 0 & \text{otherwise,} \end{cases} = \begin{cases} 1/\Delta & i\Delta \leq x < (i+1)\Delta, \\ 0 & \text{otherwise.} \end{cases} \quad (7.13)$$

Given  $B_i$ , the conditional PDF of  $X$  is uniform over the  $i$ th quantization interval.

# Theorem 7.2

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For random variable  $X$  resulting from an experiment with partition

$$B_1, \dots, B_m,$$

Discrete: 
$$P_X(x) = \sum_{i=1}^m P_{X|B_i}(x) \mathbb{P}[B_i]$$

Continuous: 
$$f_X(x) = \sum_{i=1}^m f_{X|B_i}(x) \mathbb{P}[B_i]$$

## **Proof: Theorem 7.2**

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The theorem follows directly from Theorem 1.9 with  $A = \{X = x\}$  for discrete  $X$  or  $A = \{x < X \leq x + dx\}$  when  $X$  is continuous.



## Example 7.6 Problem

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Let  $X$  denote the number of additional years that a randomly chosen 70-year-old person will live. If the person has high blood pressure, denoted as event  $H$ , then  $X$  is a geometric ( $p = 0.1$ ) random variable. Otherwise, if the person's blood pressure is normal, event  $N$ ,  $X$  has a geometric ( $p = 0.05$ ) PMF. Find the conditional PMFs  $P_{X|H}(x)$  and  $P_{X|N}(x)$ . If 40 percent of all 70-year-olds have high blood pressure, what is the PMF of  $X$ ?

## Example 7.6 Solution

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The problem statement specifies the conditional PMFs in words. Mathematically, the two conditional PMFs are

$$P_{X|H}(x) = \begin{cases} 0.1(0.9)^{x-1} & x = 1, 2, \dots, \\ 0 & \text{otherwise,} \end{cases}$$
$$P_{X|N}(x) = \begin{cases} 0.05(0.95)^{x-1} & x = 1, 2, \dots, \\ 0 & \text{otherwise.} \end{cases}$$

Since  $H, N$  is a partition, we can use Theorem 7.2 to write

$$\begin{aligned} P_X(x) &= P_{X|H}(x) P[H] + P_{X|N}(x) P[N] \\ &= \begin{cases} (0.4)(0.1)(0.9)^{x-1} + (0.6)(0.05)(0.95)^{x-1} & x = 1, 2, \dots, \\ 0 & \text{otherwise.} \end{cases} \end{aligned} \tag{7.14}$$

## Example 7.7 Problem

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Random variable  $X$  is a voltage at the receiver of a modem. When symbol “0” is transmitted (event  $B_0$ ),  $X$  is the Gaussian  $(-5, 2)$  random variable. When symbol “1” is transmitted (event  $B_1$ ),  $X$  is the Gaussian  $(5, 2)$  random variable. Given that symbols “0” and “1” are equally likely to be sent, what is the PDF of  $X$ ?

## Example 7.7 Solution

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The problem statement implies that  $P[B_0] = P[B_1] = 1/2$  and

$$f_{X|B_0}(x) = \frac{1}{2\sqrt{2\pi}}e^{-(x+5)^2/8}, \quad f_{X|B_1}(x) = \frac{1}{2\sqrt{2\pi}}e^{-(x-5)^2/8}. \quad (7.15)$$

By Theorem 7.2,

$$\begin{aligned} f_X(x) &= f_{X|B_0}(x) P[B_0] + f_{X|B_1}(x) P[B_1] \\ &= \frac{1}{4\sqrt{2\pi}} \left( e^{-(x+5)^2/8} + e^{-(x-5)^2/8} \right). \end{aligned} \quad (7.16)$$

Problem 7.7.1 asks the reader to graph  $f_X(x)$  to show its similarity to Figure 4.3.

## Quiz 7.1(A)

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On the Internet, data is transmitted in packets. In a simple model for World Wide Web traffic, the number of packets  $N$  needed to transmit a Web page depends on whether the page has graphic images. If the page has images (event  $I$ ), then  $N$  is uniformly distributed between 1 and 50 packets. If the page is just text (event  $T$ ), then  $N$  is uniform between 1 and 5 packets. Assuming a page has images with probability  $1/4$ , find the

- (a) conditional PMF  $P_{N|I}(n)$
- (b) conditional PMF  $P_{N|T}(n)$
- (c) PMF  $P_N(n)$
- (d) conditional PMF  $P_{N|N \leq 10}(n)$

## Quiz 7.1(A) Solution

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- (a) From the problem statement, we learn that the conditional PMF of  $N$  given the event  $I$  is

$$P_{N|I}(n) = \begin{cases} 0.02 & n = 1, 2, \dots, 50, \\ 0 & \text{otherwise.} \end{cases}$$

- (b) Also from the problem statement, the conditional PMF of  $N$  given the event  $T$  is

$$P_{N|T}(n) = \begin{cases} 0.2 & n = 1, \dots, 5, \\ 0 & \text{otherwise.} \end{cases}$$

- (c) The problem statement tells us that  $P[T] = 1 - P[I] = 3/4$ . From Theorem 7.2, we find the PMF of  $N$  is

$$P_N(n) = P_{N|T}(n) P[T] + P_{N|I}(n) P[I] = \begin{cases} 0.155 & n = 1, \dots, 5, \\ 0.005 & n = 6, \dots, 50, \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

[Continued]

# Quiz 7.1(A) Solution

# (Continued 2)

(d) First we find

$$P[N \leq 10] = \sum_{n=1}^{10} P_N(n) = (0.155)(5) + (0.005)(5) = 0.80. \quad (2)$$

By Theorem 7.1, the conditional PMF of  $N$  given  $N \leq 10$  is

$$\begin{aligned} P_{N|N \leq 10}(n) &= \begin{cases} \frac{P_N(n)}{P[N \leq 10]} & n \leq 10, \\ 0 & \text{otherwise,} \end{cases} \\ &= \begin{cases} \frac{0.155}{0.8} = 0.19375 & n = 1, \dots, 5, \\ \frac{0.005}{0.8} = 0.00625 & n = 6, \dots, 10, \\ 0 & \text{otherwise.} \end{cases} \end{aligned} \quad (3)$$

## Quiz 7.1(B)

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$Y$  is a continuous uniform  $(0, 10)$  random variable. Find the following:

(a)  $P[Y \leq 6]$

(b) the conditional PDF  $f_{Y|Y \leq 6}(y)$

(c)  $P[Y > 8]$

(d) the conditional PDF  $f_{Y|Y > 8}(y)$



# Quiz 7.1(B) Solution

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From the problem statement,

$$f_Y(y) = \begin{cases} 1/10 & 0 < y < 10, \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

Using this PDF and Definition 7.3, the parts are straightforward.

(a)  $P[Y \leq 6] = \int_{-\infty}^6 f_Y(y) dy = \int_0^6 (1/10) dy = 0.6.$

(b) From Definition 7.3, the conditional PDF of  $Y$  given  $Y \leq 6$  is

$$\begin{aligned} f_{Y|Y \leq 6}(y) &= \begin{cases} \frac{f_Y(y)}{P[Y \leq 6]} & y \leq 6, \\ 0 & \text{otherwise,} \end{cases} \\ &= \begin{cases} 1/6 & 0 \leq y \leq 6, \\ 0 & \text{otherwise.} \end{cases} \end{aligned} \quad (2)$$

(c) The probability  $Y > 8$  is

$$P[Y > 8] = \int_8^{10} \frac{1}{10} dy = 0.2. \quad (3)$$

(d) From Definition 7.3, the conditional PDF of  $Y$  given  $Y > 8$  is

$$\begin{aligned} f_{Y|Y > 8}(y) &= \begin{cases} \frac{f_Y(y)}{P[Y > 8]} & y > 8, \\ 0 & \text{otherwise,} \end{cases} \\ &= \begin{cases} \frac{1}{2} & 8 < y \leq 10, \\ 0 & \text{otherwise.} \end{cases} \end{aligned} \quad (4)$$

## Section 7.2

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# Conditional Expected Value Given an Event

# Theorem 7.3

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## Discrete $X$ :

- (a) For any  $x \in B$ ,  $P_{X|B}(x) \geq 0$ .
- (b)  $\sum_{x \in B} P_{X|B}(x) = 1$ .
- (c) The conditional probability that  $X$  is in the set  $C$  is

$$P[C|B] = \sum_{x \in C} P_{X|B}(x).$$

## Continuous $X$ :

- (a) For any  $x \in B$ ,  $f_{X|B}(x) \geq 0$ .
- (b)  $\int_B f_{X|B}(x) dx = 1$ .
- (c) The conditional probability that  $X$  is in the set  $C$  is

$$P[C|B] = \int_C f_{X|B}(x) dx.$$

## Definition 7.4 Conditional Expected Value

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The conditional expected value of random variable  $X$  given condition  $B$  is

$$\text{Discrete: } E[X|B] = \sum_{x \in B} x P_{X|B}(x)$$

$$\text{Continuous: } E[X|B] = \int_{-\infty}^{\infty} x f_{X|B}(x) dx.$$

## Theorem 7.4

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For a random variable  $X$  resulting from an experiment with partition  $B_1, \dots, B_m$ ,

$$E[X] = \sum_{i=1}^m E[X|B_i] P[B_i].$$

## Proof: Theorem 7.4

When  $X$  is discrete,  $E[X] = \sum_x x P_X(x)$ , and we can use Theorem 7.2 to write

$$\begin{aligned} E[X] &= \sum_x x \sum_{i=1}^m P_{X|B_i}(x) P[B_i] \\ &= \sum_{i=1}^m P[B_i] \sum_x x P_{X|B_i}(x) = \sum_{i=1}^m P[B_i] E[X|B_i]. \end{aligned} \quad (7.17)$$

When  $X$  is continuous, the proof uses the continuous version of Theorem 7.2 and follows the same logic, with the summation over  $x$  replaced by integration.

## Theorem 7.5

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The conditional expected value of  $Y = g(X)$  given condition  $B$  is

$$\text{Discrete: } E[Y|B] = E[g(X)|B] = \sum_{x \in B} g(x) P_{X|B}(x)$$

$$\text{Continuous: } E[Y|B] = E[g(X)|B] = \int_{-\infty}^{\infty} g(x) f_{X|B}(x) dx.$$

# Conditional Variance and

## **Definition 7.5 Standard Deviation**

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*The conditional variance of  $X$  given event  $B$  is*

$$\text{Var}[X|B] = \text{E} \left[ (X - \mu_{X|B})^2 | B \right] = \text{E} [X^2|B] - \mu_{X|B}^2.$$

*The conditional standard deviation is  $\sigma_{X|B} = \sqrt{\text{Var}[X|B]}$ .*



## Example 7.8 Problem

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Find the conditional expected value, the conditional variance, and the conditional standard deviation for the long videos defined in Example 7.2.

## Example 7.8 Solution

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$$E[X|L] = \mu_{X|L} = \sum_{x=5}^8 x P_{X|L}(x) = 0.25 \sum_{x=5}^8 x = 6.5 \text{ minutes.} \quad (7.18)$$

$$E[X^2|L] = 0.25 \sum_{x=5}^8 x^2 = 43.5 \text{ minutes}^2. \quad (7.19)$$

$$\text{Var}[X|L] = E[X^2|L] - \mu_{X|L}^2 = 1.25 \text{ minutes}^2. \quad (7.20)$$

$$\sigma_{X|L} = \sqrt{\text{Var}[X|L]} = 1.12 \text{ minutes.} \quad (7.21)$$

## Quiz 7.2(A)

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Continuing Quiz 7.1(A), find

(a)  $E[N|N \leq 10]$ ,

(b)  $\text{Var}[N|N \leq 10]$ .

## Quiz 7.2(A) Solution

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We refer to the solution of Quiz 7.1(A) for  $P_{N|N \leq 10}(n)$ .

(a) Given  $P_{N|N \leq 10}(n)$ , calculating a conditional expected value is the same as for any other expected value except we use the conditional PMF.

$$\begin{aligned} E[N|N \leq 10] &= \sum_n n P_{N|N \leq 10}(n) \\ &= \sum_{n=1}^5 0.19375n + \sum_{n=6}^{10} 0.00625n = 3.15625. \end{aligned} \quad (1)$$

(b) For the conditional variance, we first find the conditional second moment

$$\begin{aligned} E[N^2|N \leq 10] &= \sum_n n^2 P_{N|N \leq 10}(n) \\ &= \sum_{n=1}^5 0.19375n^2 + \sum_{n=6}^{10} 0.00625n^2 \\ &= 0.19375(55) + 0.00625(330) = 12.719. \end{aligned} \quad (2)$$

The conditional variance is

$$\begin{aligned} \text{Var}[N|N \leq 10] &= E[N^2|N \leq 10] - (E[N|N \leq 10])^2 \\ &= 12.719 - (3.156)^2 = 2.757. \end{aligned} \quad (3)$$

## Quiz 7.2(B)

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Continuing Quiz 7.1(B), find

(a)  $E[Y|Y \leq 6]$ ,

(b)  $\text{Var}[Y|Y \leq 6]$ .

## Quiz 7.2(B) Solution

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We refer to the solution of Quiz 7.1(B) for the conditional PDFs  $f_{Y|Y \leq 6}(y)$  and  $f_{Y|Y > 8}(y)$ .

(a) From  $f_{Y|Y \leq 6}(y)$ , the conditional expectation is

$$E[Y|Y \leq 6] = \int_{-\infty}^{\infty} y f_{Y|Y \leq 6}(y) dy = \int_0^6 \frac{y}{6} dy = 3. \quad (1)$$

(b) From the conditional PDF  $f_{Y|Y > 8}(y)$ , we see that given  $Y > 8$ ,  $Y$  is conditionally a continuous uniform ( $a = 8, b = 10$ ) random variable. Thus,

$$\text{Var}[Y|Y > 8] = (b - a)^2 / 12 = 1/3. \quad (2)$$

## Section 7.3

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# Conditioning Two Random Variables by an Event

## **Definition 7.6 Conditional Joint PMF**

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*For discrete random variables  $X$  and  $Y$  and an event  $B$  with  $P[B] > 0$ , the conditional joint PMF of  $X$  and  $Y$  given  $B$  is*

$$P_{X,Y|B}(x, y) = P[X = x, Y = y|B].$$



## Theorem 7.6

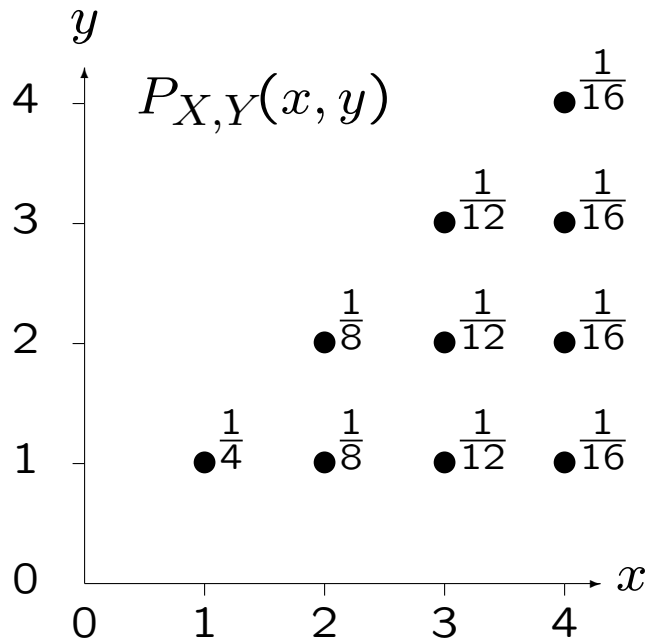
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For any event  $B$ , a region of the  $X, Y$  plane with  $P[B] > 0$ ,

$$P_{X,Y|B}(x, y) = \begin{cases} \frac{P_{X,Y}(x, y)}{P[B]} & (x, y) \in B, \\ 0 & \text{otherwise.} \end{cases}$$

## Example 7.9 Problem

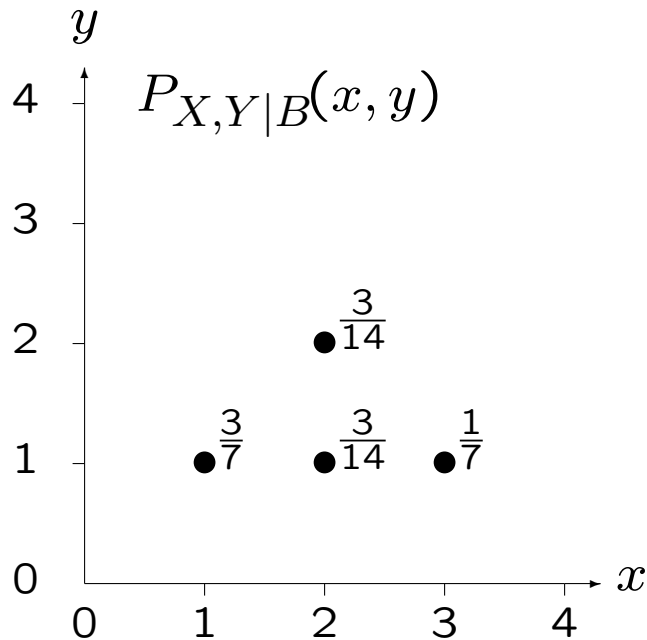
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Random variables  $X$  and  $Y$  have the joint PMF  $P_{X,Y}(x,y)$  as shown. Let  $B = \{X + Y \leq 4\}$  and find the conditional PMF  $P_{X,Y|B}(x,y)$ .

## Example 7.9 Solution

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Event  $B = \{(1, 1), (2, 1), (2, 2), (3, 1)\}$  consists of all points  $(x, y)$  such that  $x + y \leq 4$ . By adding up the probabilities of all outcomes in  $B$ , we find

$$\begin{aligned} P[B] &= P_{X,Y}(1, 1) + P_{X,Y}(2, 1) \\ &\quad + P_{X,Y}(2, 2) + P_{X,Y}(3, 1) = \frac{7}{12}. \end{aligned}$$

The conditional PMF  $P_{X,Y|B}(x, y)$  is shown on the left.

## **Definition 7.7 Conditional Joint PDF**

*Given an event  $B$  with  $P[B] > 0$ , the conditional joint probability density function of  $X$  and  $Y$  is*

$$f_{X,Y|B}(x,y) = \begin{cases} \frac{f_{X,Y}(x,y)}{P[B]} & (x,y) \in B, \\ 0 & \text{otherwise.} \end{cases}$$

## Example 7.10 Problem

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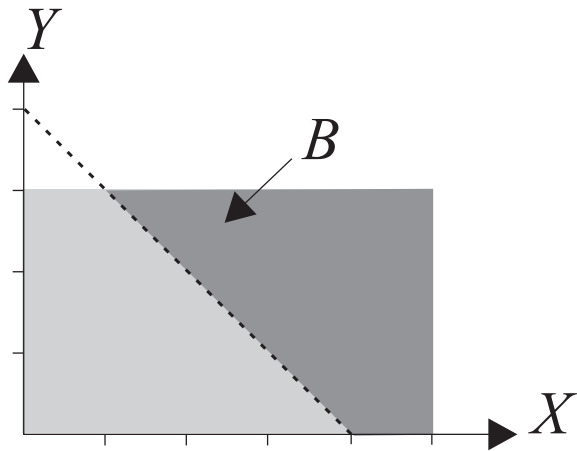
$X$  and  $Y$  are random variables with joint PDF

$$f_{X,Y}(x,y) = \begin{cases} 1/15 & 0 \leq x \leq 5, 0 \leq y \leq 3, \\ 0 & \text{otherwise.} \end{cases} \quad (7.22)$$

Find the conditional PDF of  $X$  and  $Y$  given the event  $B = \{X + Y \geq 4\}$ .

## Example 7.10 Solution

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We calculate  $P[B]$  by integrating  $f_{X,Y}(x,y)$  over the region  $B$ .

$$\begin{aligned} P[B] &= \int_0^3 \int_{4-y}^5 \frac{1}{15} dx dy \\ &= \frac{1}{15} \int_0^3 (1+y) dy = 1/2. \end{aligned} \quad (7.23)$$

Definition 7.7 leads to the conditional joint PDF

$$f_{X,Y|B}(x,y) = \begin{cases} 2/15 & 0 \leq x \leq 5, 0 \leq y \leq 3, x+y \geq 4, \\ 0 & \text{otherwise.} \end{cases} \quad (7.24)$$

## Theorem 7.7      Conditional Expected Value

For random variables  $X$  and  $Y$  and an event  $B$  of nonzero probability, the conditional expected value of  $W = g(X, Y)$  given  $B$  is

$$\text{Discrete: } E[W|B] = \sum_{x \in S_X} \sum_{y \in S_Y} g(x, y) P_{X,Y|B}(x, y)$$

$$\text{Continuous: } E[W|B] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y|B}(x, y) dx dy.$$

## Example 7.11 Problem

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Continuing Example 7.9, find the conditional expected value and the conditional variance of  $W = X + Y$  given the event  $B = \{X + Y \leq 4\}$ .



## Example 7.11 Solution

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We recall from Example 7.9 that  $P_{X,Y|B}(x,y)$  has four points with nonzero probability:  $(1,1)$ ,  $(1,2)$ ,  $(1,3)$ , and  $(2,2)$ . Their probabilities are  $3/7$ ,  $3/14$ ,  $1/7$ , and  $3/14$ , respectively. Therefore,

$$\begin{aligned} \mathbb{E}[W|B] &= \sum_{x,y} (x+y) P_{X,Y|B}(x,y) \\ &= 2 \left(\frac{3}{7}\right) + 3 \left(\frac{3}{14}\right) + 4 \left(\frac{1}{7}\right) + 4 \left(\frac{3}{14}\right) = \frac{41}{14}. \end{aligned} \quad (7.25)$$

Similarly,

$$\begin{aligned} \mathbb{E}[W^2|B] &= \sum_{x,y} (x+y)^2 P_{X,Y|B}(x,y) \\ &= 2^2 \left(\frac{3}{7}\right) + 3^2 \left(\frac{3}{14}\right) + 4^2 \left(\frac{1}{7}\right) + 4^2 \left(\frac{3}{14}\right) = \frac{131}{14}. \end{aligned} \quad (7.26)$$

The conditional variance is

$$\text{Var}[W|B] = \mathbb{E}[W^2|B] - (\mathbb{E}[W|B])^2 = \frac{131}{14} - \left(\frac{41}{14}\right)^2 = \frac{153}{196}. \quad (7.27)$$

## Quiz 7.3(A)

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Random variables  $L$  and  $X$  have joint PMF

$P_{L,X}(l, x)$	$x = 40$	$x = 60$
$l = 1$	0.15	0.1
$l = 2$	0.3	0.2
$l = 3$	0.15	0.1

(7.28)

For random variable  $V = LX$ , we define the event  $A = \{V > 80\}$ . Find the conditional PMF  $P_{L,X|A}(l, x)$ . What are  $E[V|A]$  and  $\text{Var}[V|A]$ ?

## Quiz 7.3(A) Solution

---

Since the event  $V > 80$  occurs only for the pairs  $(L, X) = (2, 60)$ ,  $(L, X) = (3, 40)$  and  $(L, X) = (3, 60)$ ,

$$\begin{aligned} P[A] &= P[V > 80] = P_{L,X}(2, 60) + P_{L,X}(3, 40) + P_{L,X}(3, 60) \\ &= 0.45. \end{aligned} \tag{1}$$

By Definition 7.6,

$$P_{L,X|A}(l, X) = \begin{cases} \frac{P_{L,X}(l, x)}{P[A]} & lx > 80, \\ 0 & \text{otherwise.} \end{cases}$$

We can represent this conditional PMF in the following table:

$P_{L,X A}(l, x)$	$x = 40$	$x = 60$
$l = 1$	0	0
$l = 2$	0	4/9
$l = 3$	1/3	2/9

[Continued]

## Quiz 7.3(A) Solution

## (Continued 2)

The conditional expectation of  $V$  can be found from the conditional PMF.

$$\begin{aligned} E[V|A] &= \sum_l \sum_x lx P_{L,X|A}(l, x) \\ &= (120)\frac{4}{9} + (120)\frac{1}{3} + (180)\frac{2}{9} = 133\frac{1}{3}. \end{aligned} \quad (2)$$

For the conditional variance  $\text{Var}[V|A]$ , we first find the conditional second moment

$$\begin{aligned} E[V^2|A] &= \sum_l \sum_x (lx)^2 P_{L,X|A}(l, x) \\ &= (120)^2\frac{4}{9} + (120)^2\frac{1}{3} + (180)^2\frac{2}{9} = 18,400. \end{aligned} \quad (3)$$

It follows that

$$\text{Var}[V|A] = E[V^2|A] - (E[V|A])^2 = 622\frac{2}{9} \quad (4)$$

## Quiz 7.3(B)

---

Random variables  $X$  and  $Y$  have the joint PDF

$$f_{X,Y}(x,y) = \begin{cases} xy/4000 & 1 \leq x \leq 3, 40 \leq y \leq 60, \\ 0 & \text{otherwise.} \end{cases} \quad (7.29)$$

For random variable  $W = XY$ , we define the event  $B = \{W > 80\}$ . Find the conditional joint PDF  $f_{X,Y|B}(x,y)$ . What are  $E[W|B]$  and  $\text{Var}[W|B]$ ?

## Quiz 7.3(B) Solution

---

For continuous random variables  $X$  and  $Y$ , we first calculate the probability of the conditioning event.

$$P[B] = \iint_B f_{X,Y}(x,y) dx dy = \int_{40}^{60} \int_{80/y}^3 \frac{xy}{4000} dx dy. \quad (1)$$

A little calculus yields

$$\begin{aligned} P[B] &= \int_{40}^{60} \frac{y}{4000} \left( \frac{x^2}{2} \Big|_{80/y}^3 \right) dy \\ &= \int_{40}^{60} \frac{y}{4000} \left( \frac{9}{2} - \frac{3200}{y^2} \right) dy = \frac{9}{8} - \frac{4}{5} \ln \frac{3}{2}. \end{aligned} \quad (2)$$

In fact,  $P[B] \approx 0.801$ . The conditional PDF of  $X$  and  $Y$  is

$$f_{X,Y|B}(x,y) = \begin{cases} \frac{f_{X,Y}(x,y)}{P[B]} & (x,y) \in B, \\ 0 & \text{otherwise,} \end{cases} = \begin{cases} Kxy & 40 \leq y \leq 60, \\ & 80/y \leq x \leq 3, \\ 0 & \text{otherwise.} \end{cases} \quad (3)$$

where  $K = (4000 P[B])^{-1}$ . The conditional expectation of  $W$  given event  $B$  is [\[Continued\]](#)

## Quiz 7.3(B) Solution

## (Continued 2)

$$E[W|B] = \iint xy f_{X,Y|B}(x,y) dx dy = \int_{40}^{60} \int_{80/y}^3 Kx^2y^2 dx dy. \quad (4)$$

These next steps are just calculus:

$$\begin{aligned} E[W|B] &= \frac{K}{3} \int_{40}^{60} y^2 x^3 \Big|_{x=80/y}^{x=3} dy \\ &= \frac{K}{3} \int_{40}^{60} (27y^2 - 80^3/y) dy = \frac{K}{3} (9y^3 - 80^3 \ln y) \Big|_{40}^{60} \approx 120.78. \end{aligned} \quad (5)$$

The conditional second moment of  $K$  given  $B$  is

$$E[W^2|B] = \iint (xy)^2 f_{X,Y|B}(x,y) dx dy = \int_{40}^{60} \int_{80/y}^3 Kx^3y^3 dx dy. \quad (6)$$

With a final bit of calculus,

$$\begin{aligned} E[W^2|B] &= \frac{K}{4} \int_{40}^{60} y^3 x^4 \Big|_{x=80/y}^{x=3} dy \\ &= \frac{K}{4} \int_{40}^{60} (81y^3 - 80^4/y) dy = \frac{K}{4} \left( \frac{81}{4} y^4 - 80^4 \ln y \right) \Big|_{40}^{60} \\ &\approx 16,116.10. \end{aligned} \quad (7)$$

It follows that  $\text{Var}[W|B] = E[W^2|B] - (E[W|B])^2 \approx 1528.30$ .

## Section 7.4

---

# Conditioning by a Random Variable



## **Definition 7.8 Conditional PMF**

---

*For any event  $Y = y$  such that  $P_Y(y) > 0$ , the conditional PMF of  $X$  given  $Y = y$  is*

$$P_{X|Y}(x|y) = P[X = x|Y = y].$$

## Theorem 7.8

---

For discrete random variables  $X$  and  $Y$  with joint PMF  $P_{X,Y}(x, y)$ , and  $x$  and  $y$  such that  $P_X(x) > 0$  and  $P_Y(y) > 0$ ,

$$P_{X|Y}(x|y) = \frac{P_{X,Y}(x, y)}{P_Y(y)}, \quad P_{Y|X}(y|x) = \frac{P_{X,Y}(x, y)}{P_X(x)}.$$

## **Proof: Theorem 7.8**

---

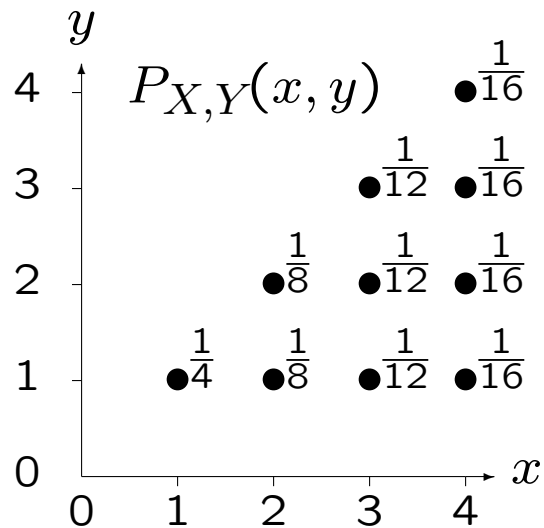
Referring to Definition 7.8, Definition 1.5, and Theorem 5.4, we observe that

$$P_{X|Y}(x|y) = P[X = x|Y = y] = \frac{P[X = x, Y = y]}{P[Y = y]} = \frac{P_{X,Y}(x, y)}{P_Y(y)}. \quad (7.30)$$

The proof of the second part is the same with  $X$  and  $Y$  reversed.

## Example 7.12 Problem

---



Random variables  $X$  and  $Y$  have the joint PMF  $P_{X,Y}(x,y)$ , as given in Example 7.9 and repeated in the accompanying graph. Find the conditional PMF of  $Y$  given  $X = x$  for each  $x \in S_X$ .

## Example 7.12 Solution

---

To apply Theorem 7.8, we first find the marginal PMF  $P_X(x)$ . By Theorem 5.4,  $P_X(x) = \sum_{y \in S_Y} P_{X,Y}(x, y)$ . For a given  $X = x$ , we sum the nonzero probabilities along the vertical line  $X = x$ . That is,

$$P_X(x) = \begin{cases} 1/4 & x = 1, \\ 1/8 + 1/8 & x = 2, \\ 1/12 + 1/12 + 1/12 & x = 3, \\ 1/16 + 1/16 + 1/16 + 1/16 & x = 4, \\ 0 & \text{otherwise,} \end{cases} = \begin{cases} 1/4 & x = 1, \\ 1/4 & x = 2, \\ 1/4 & x = 3, \\ 1/4 & x = 4, \\ 0 & \text{otherwise.} \end{cases} \quad (7.31)$$

Theorem 7.8 implies that for  $x \in \{1, 2, 3, 4\}$ ,

$$P_{Y|X}(y|x) = \frac{P_{X,Y}(x, y)}{P_X(x)} = 4P_{X,Y}(x, y). \quad (7.32)$$

[Continued]

## Example 7.12 Solution

(Continued 2)

---

For each  $x \in \{1, 2, 3, 4\}$ ,  $P_{Y|X}(y|x)$  is a different PMF.

$$P_{Y|X}(y|1) = \begin{cases} 1 & y = 1, \\ 0 & \text{otherwise.} \end{cases}$$

$$P_{Y|X}(y|2) = \begin{cases} 1/2 & y \in \{1, 2\}, \\ 0 & \text{otherwise.} \end{cases}$$

$$P_{Y|X}(y|3) = \begin{cases} 1/3 & y \in \{1, 2, 3\}, \\ 0 & \text{otherwise.} \end{cases}$$

$$P_{Y|X}(y|4) = \begin{cases} 1/4 & y \in \{1, 2, 3, 4\}, \\ 0 & \text{otherwise.} \end{cases}$$

Given  $X = x$ ,  $Y$  is conditionally a discrete uniform  $(1, x)$  random variable.

## **Definition 7.9 Conditional PDF**

---

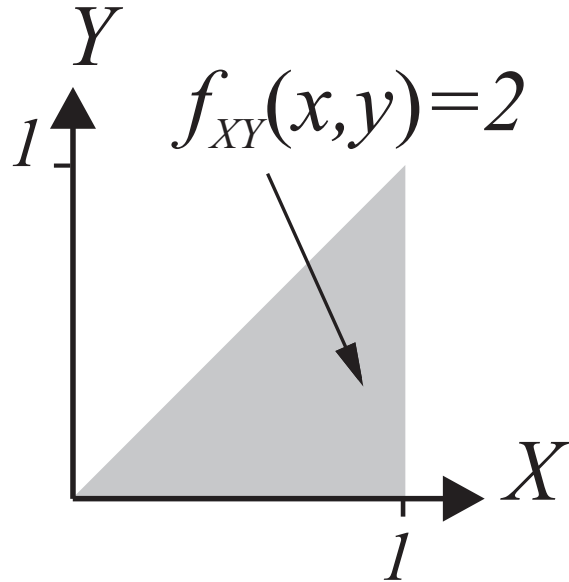
For  $y$  such that  $f_Y(y) > 0$ , the conditional PDF of  $X$  given  $\{Y = y\}$  is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}.$$

## Example 7.13 Problem

---

Returning to Example 5.8, random variables  $X$  and  $Y$  have joint PDF



$$f_{X,Y}(x, y) = \begin{cases} 2 & 0 \leq y \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.34)$$

For  $0 \leq x \leq 1$ , find the conditional PDF  $f_{Y|X}(y|x)$ . For  $0 \leq y \leq 1$ , find the conditional PDF  $f_{X|Y}(x|y)$ .



## Example 7.13 Solution

---

For  $0 \leq x \leq 1$ , Theorem 5.8 implies

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy = \int_0^x 2 dy = 2x. \quad (7.35)$$

The conditional PDF of  $Y$  given  $X$  is

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x, y)}{f_X(x)} = \begin{cases} 1/x & 0 \leq y \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.36)$$

Given  $X = x$ , we see that  $Y$  is the uniform  $(0, x)$  random variable. For  $0 \leq y \leq 1$ , Theorem 5.8 implies

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx = \int_y^1 2 dx = 2(1 - y). \quad (7.37)$$

Furthermore, Equation (7.33) implies

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)} = \begin{cases} 1/(1 - y) & y \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.38)$$

Conditioned on  $Y = y$ , we see that  $X$  is the uniform  $(y, 1)$  random variable.

## **Theorem 7.9**

---

For discrete random variables  $X$  and  $Y$  with joint PMF  $P_{X,Y}(x, y)$ , and  $x$  and  $y$  such that  $P_X(x) > 0$  and  $P_Y(y) > 0$ ,

$$P_{X,Y}(x, y) = P_{Y|X}(y|x) P_X(x) = P_{X|Y}(x|y) P_Y(y).$$

## **Theorem 7.10**

---

For continuous random variables  $X$  and  $Y$  with joint PDF  $f_{X,Y}(x, y)$ , and  $x$  and  $y$  such that  $f_X(x) > 0$  and  $f_Y(y) > 0$ ,

$$f_{X,Y}(x, y) = f_{Y|X}(y|x) f_X(x) = f_{X|Y}(x|y) f_Y(y).$$

## Example 7.14 Problem

---

Let  $R$  be the uniform  $(0, 1)$  random variable. Given  $R = r$ ,  $X$  is the uniform  $(0, r)$  random variable. Find the conditional PDF of  $R$  given  $X$ .

## Example 7.14 Solution

---

The problem definition states that

$$f_R(r) = \begin{cases} 1 & 0 \leq r < 1, \\ 0 & \text{otherwise,} \end{cases} \quad f_{X|R}(x|r) = \begin{cases} 1/r & 0 \leq x < r, \\ 0 & \text{otherwise.} \end{cases} \quad (7.39)$$

It follows from Theorem 7.10 that the joint PDF of  $R$  and  $X$  is

$$f_{R,X}(r, x) = f_{X|R}(x|r) f_R(r) = \begin{cases} 1/r & 0 \leq x < r < 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.40)$$

Now we can find the marginal PDF of  $X$  from Theorem 5.8. For  $0 < x < 1$ ,

$$f_X(x) = \int_{-\infty}^{\infty} f_{R,X}(r, x) dr = \int_x^1 \frac{dr}{r} = -\ln x. \quad (7.41)$$

By the definition of the conditional PDF,

$$f_{R|X}(r|x) = \frac{f_{R,X}(r, x)}{f_X(x)} = \begin{cases} \frac{1}{-r \ln x} & x \leq r \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.42)$$

## Definition 5.4

---

- We recall from Definition 5.4 that discrete random variables  $X$  and  $Y$  are independent if  $P_{X,Y}(x, y) = P_X(x)P_Y(y)$ .
- Similarly, continuous  $X$  and  $Y$  are independent if  $f_{X,Y}(x, y) = f_X(x)f_Y(y)$ .
- We apply these observations to Theorems 7.9 and 7.10.

# Theorem 7.11

---

If  $X$  and  $Y$  are independent,

Discrete:  $P_{X|Y}(x|y) = P_X(x),$   $P_{Y|X}(y|x) = P_Y(y);$

Continuous:  $f_{X|Y}(x|y) = f_X(x),$   $f_{Y|X}(y|x) = f_Y(y).$

## Quiz 7.4(A)

---

The PMF of random variable  $X$  satisfies  $P_X(0) = 1 - P_X(2) = 0.4$ . The conditional probability model for random variable  $Y$  given  $X$  is

$$P_{Y|X}(y|0) = \begin{cases} 0.8 & y = 0, \\ 0.2 & y = 1, \\ 0 & \text{otherwise,} \end{cases} \quad P_{Y|X}(y|2) = \begin{cases} 0.5 & y = 0, \\ 0.5 & y = 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.43)$$

- (a) What is the probability model for  $X$  and  $Y$ ? Write the joint PMF  $P_{X,Y}(x,y)$  as a table.
- (b) If  $Y = 0$ , what is the conditional PMF  $P_{X|Y}(x|0)$ ?



## Quiz 7.4(A) Solution

- (a) The joint PMF of  $X$  and  $Y$  can be found from the marginal and conditional PMFs via  $P_{X,Y}(x,y) = P_{Y|X}(y|x)P_X(x)$ . Incorporating the information from the given conditional PMFs can be confusing, however. Consequently, we note that  $X$  has range  $S_X = \{0, 2\}$  and  $Y$  has range  $S_Y = \{0, 1\}$ . A table of the joint PMF will include all four possible combinations of  $X$  and  $Y$ . The general form of the table is

$P_{X,Y}(x,y)$	$y = 0$	$y = 1$
$x = 0$	$P_{Y X}(0 0)P_X(0)$	$P_{Y X}(1 0)P_X(0)$
$x = 2$	$P_{Y X}(0 2)P_X(2)$	$P_{Y X}(1 2)P_X(2)$

Substituting values from  $P_{Y|X}(y|x)$  and  $P_X(x)$ , we have

$P_{X,Y}(x,y)$	$y = 0$	$y = 1$
$x = 0$	$(0.8)(0.4)$	$(0.2)(0.4)$
$x = 2$	$(0.5)(0.6)$	$(0.5)(0.6)$

which simplifies to

$P_{X,Y}(x,y)$	$y = 0$	$y = 1$
$x = 0$	0.32	0.08
$x = 2$	0.3	0.3

- (b) From the joint PMF  $P_{X,Y}(x,y)$ , we can calculate  $P_Y(0) = 0.32 + 0.3 = 0.62$  and the conditional PMF

$$P_{X|Y}(x|0) = \frac{P_{X,Y}(x,0)}{P_Y(0)} = \begin{cases} \frac{0.32}{0.62} = \frac{16}{31} & x = 0, \\ \frac{0.3}{0.62} = \frac{15}{31} & x = 2, \\ 0 & \text{otherwise.} \end{cases}$$

## Quiz 7.4(B)

---

The PDF of  $Y$  given  $X$  are

$$f_X(x) = \begin{cases} 3x^2 & 0 < x \leq 1, \\ 0 & \text{otherwise,} \end{cases} \quad f_{Y|X}(y|x) = \begin{cases} 2y/x^2 & 0 \leq y \leq x, \\ 0 & \text{otherwise.} \end{cases} \quad (7.44)$$

(a) What is the probability model for  $X$  and  $Y$ ? Find  $f_{X,Y}(x,y)$ .

(b) If  $Y = 1/2$ , what is the conditional PDF  $f_{X|Y}(x|1/2)$ ?

## Quiz 7.4(B) Solution

---

(a) The joint PDF of  $X$  and  $Y$  is

$$f_{X,Y}(x,y) = f_{Y|X}(y|x) f_X(x) = \begin{cases} 6y & 0 \leq y \leq x, 0 \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases}$$

(b) To find  $f_{X|Y}(x|1/2)$ , we first find

$$f_Y(1/2) = \int_{-\infty}^{\infty} f_{X,Y}(x, 1/2) dx.$$

For this integral, we keep in mind that  $f_{X,Y}(x,y)$  is nonzero for  $y \leq x \leq 1$ . Specifically, for  $y = 1/2$ , we integrate over  $1/2 \leq x \leq 1$ :

$$f_Y(1/2) = \int_{1/2}^1 6(1/2) dx = 3/2. \quad (1)$$

For  $1/2 \leq x \leq 1$ , the conditional PDF of  $X$  given  $Y = 1/2$  is

$$f_{X|Y}(x|1/2) = \frac{f_{X,Y}(x, 1/2)}{f_Y(1/2)} = \frac{6(1/2)}{3/2} = 2. \quad (2)$$

For  $x < 1/2$  or  $x > 1$ ,  $f_{X|Y}(x|1/2) = 0$ . Thus given  $Y = 1/2$ , the  $X$  has the continuous uniform  $(1/2, 1)$  PDF

$$f_{X|Y}(x|1/2) = \begin{cases} 2 & \frac{1}{2} \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (3)$$

## Section 7.5

---

# Conditional Expected Value Given a Random Variable

# Conditional Expected Value

## **Definition 7.10** of a Function

---

For any  $y \in S_Y$ , the conditional expected value of  $g(X, Y)$  given  $Y = y$  is

$$\text{Discrete: } E[g(X, Y)|Y = y] = \sum_{x \in S_X} g(x, y)P_{X|Y}(x|y)$$

$$\text{Continuous: } E[g(X, Y)|Y = y] = \int_{-\infty}^{\infty} g(x, y)f_{X|Y}(x|y) dx.$$

## Example 7.15 Problem

---

In Example 7.12, we derived conditional PMFs  $P_{Y|X}(y|1)$ ,  $P_{Y|X}(y|2)$ ,  $P_{Y|X}(y|3)$ , and  $P_{Y|X}(y|4)$ . Find  $E[Y|X = x]$  for  $x = 1, 2, 3, 4$ .

## Example 7.15 Solution

---

In Example 7.12 we found that given  $X = x$ ,  $Y$  was a discrete uniform  $(1, x)$  random variable. Since a discrete uniform  $(1, x)$  random variable has expected value  $(1 + x)/2$ ,

$$E[Y|X = 1] = \frac{1 + 1}{2} = 1, \quad E[Y|X = 2] = \frac{1 + 2}{2} = 1.5, \quad (7.45)$$

$$E[Y|X = 3] = \frac{1 + 3}{2} = 2, \quad E[Y|X = 4] = \frac{1 + 4}{2} = 2.5. \quad (7.46)$$

# Theorem 7.12

---

For independent random variables  $X$  and  $Y$ ,

(a)  $E[X|Y = y] = E[X]$  for all  $y \in S_Y$ ,

(b)  $E[Y|X = x] = E[Y]$  for all  $x \in S_X$ .



# Proof: Theorem 7.12

---

We present the proof for discrete random variables. By replacing PMFs and sums with PDFs and integrals, we arrive at essentially the same proof for continuous random variables. Since  $P_{X|Y}(x|y) = P_X(x)$ ,

$$E[X|Y = y] = \sum_{x \in S_X} x P_{X|Y}(x|y) = \sum_{x \in S_X} x P_X(x) = E[X]. \quad (7.47)$$

Since  $P_{Y|X}(y|x) = P_Y(y)$ ,

$$E[Y|X = x] = \sum_{y \in S_Y} y P_{Y|X}(y|x) = \sum_{y \in S_Y} y P_Y(y) = E[Y]. \quad (7.48)$$

# Conditional Expected Values

---

- When we introduced the concept of expected value in Chapters 3 and 4, we observed that  $E[X]$  is a property of the probability model of  $X$ . This is also true for  $E[X|B]$  when  $P[B] > 0$ .
- The situation is more complex when we consider  $E[X|Y = y]$ , the conditional expected value given a random variable.
- In this case, the conditional expected value is a different number for each possible observation  $y \in S_Y$ .
- This implies that  $E[X|Y = y]$  is a function of the random variable  $Y$ .
- We use the notation  $E[X|Y]$  to denote this function of the random variable  $Y$ .
- Since a function of a random variable is another random variable, we conclude that  $E[X|Y]$  *is a random variable!*

# Conditional Expected Value

## **Definition 7.11** Function

---

*The conditional expected value  $E[X|Y]$  is a function of random variable  $Y$  such that if  $Y = y$ , then  $E[X|Y] = E[X|Y = y]$ .*

## Example 7.16 Problem

---

For random variables  $X$  and  $Y$  in Example 5.8, we found in Example 7.13 that the conditional PDF of  $X$  given  $Y$  is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \begin{cases} 1/(1-y) & 0 \leq y \leq x \leq 1, \\ 0 & \text{otherwise.} \end{cases} \quad (7.49)$$

Find the conditional expected values  $E[X|Y = y]$  and  $E[X|Y]$ .

## Example 7.16 Solution

---

Given the conditional PDF  $f_{X|Y}(x|y)$ , we perform the integration

$$\begin{aligned} E[X|Y = y] &= \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx \\ &= \int_y^1 \frac{1}{1-y} x dx = \frac{x^2}{2(1-y)} \Big|_{x=y}^{x=1} = \frac{1+y}{2}. \end{aligned} \quad (7.50)$$

Since  $E[X|Y = y] = (1 + y)/2$ ,  $E[X|Y] = (1 + Y)/2$ .

# Iterated Expectation Property

---

- An interesting property of the random variable  $E[X|Y]$  is its expected value  $E[E[X|Y]]$ .
- We find  $E[E[X|Y]]$  in two steps: First we calculate  $g(y) = E[X|Y = y]$ , and then we apply Theorem 4.4 to evaluate  $E[g(Y)]$ .
- This two-step process is known as *iterated expectation*.

## Theorem 7.13    Iterated Expectation

---

$$E[E[X|Y]] = E[X].$$

# Proof: Theorem 7.13

---

We consider continuous random variables  $X$  and  $Y$  and apply Theorem 4.4:

$$E[E[X|Y]] = \int_{-\infty}^{\infty} E[X|Y = y] f_Y(y) dy. \quad (7.51)$$

To obtain this formula from Theorem 4.4, we have used  $E[X|Y = y]$  in place of  $g(x)$  and  $f_Y(y)$  in place of  $f_X(x)$ . Next, we substitute the right side of Equation (7.45) for  $E[X|Y = y]$ :

$$E[E[X|Y]] = \int_{-\infty}^{\infty} \left( \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx \right) f_Y(y) dy. \quad (7.52)$$

Rearranging terms in the double integral and reversing the order of integration, we obtain

$$E[E[X|Y]] = \int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} f_{X|Y}(x|y) f_Y(y) dy dx. \quad (7.53)$$

Next, we apply Theorem 7.10 and Theorem 5.8 to infer that the inner integral is  $f_X(x)$ . Therefore,

$$E[E[X|Y]] = \int_{-\infty}^{\infty} x f_X(x) dx. \quad (7.54)$$

The proof is complete because the right side of this formula is the definition of  $E[X]$ . A similar derivation (using sums instead of integrals) proves the theorem for discrete random variables.



## Theorem 7.14

---

$$E [E [g(X)|Y]] = E [g(X)] .$$

## Quiz 7.5(A)

---

For random variables  $A$  and  $B$  in Quiz 7.4(A) find:

- (a)  $E[Y|X = 2]$ ,
- (b)  $\text{Var}[X|Y = 0]$ .

## Quiz 7.5(A) Solution

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- (a) Given the conditional PMF  $P_{Y|X}(y|2)$ , it is easy to calculate the conditional expectation

$$E[Y|X = 2] = \sum_{y=0}^1 yP_{Y|X}(y|2) = (0)(0.5) + (1)(0.5) = 0.5. \quad (1)$$

- (b) We can calculate the conditional variance  $\text{Var}[X|Y = 0]$  using the conditional PMF  $P_{X|Y}(x|0)$ . First we calculate the conditional expected value

$$E[X|Y = 0] = \sum_x xP_{X|Y}(x|0) = 0 \cdot \frac{16}{31} + 2 \cdot \frac{15}{31} = \frac{30}{31}. \quad (2)$$

The conditional second moment is

$$E[X^2|Y = 0] = \sum_x x^2P_{X|Y}(x|0) = 0^2 \frac{16}{31} + 2^2 \frac{15}{31} = \frac{60}{31}. \quad (3)$$

The conditional variance is then

$$\text{Var}[X|Y = 0] = E[X^2|Y = 0] - (E[X|Y = 0])^2 = 960/961. \quad (4)$$

## Quiz 7.5(B)

---

For random variables  $X$  and  $Y$  in Quiz 7.4(B), find:

(a)  $E[Y|X = 1/2]$ ,

(b)  $\text{Var}[X|Y = 1/2]$ .

## Quiz 7.5(B) Solution

---

(a) From the conditional PDF  $f_{Y|X}(y|x)$  given in Quiz 7.4(B),

$$f_{Y|X}(y|1/2) = \begin{cases} 8y & 0 \leq y \leq 1/2, \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

Now we calculate the conditional expected value

$$E[Y|X = 1/2] = \int_0^{1/2} y(8y) dy = 8y^3/3 \Big|_0^{1/2} = 1/3. \quad (2)$$

(b) From the solution to Quiz 7.4(B), we see that given  $Y = 1/2$ , the conditional PDF of  $X$  is uniform  $(1/2, 1)$ . Thus, by the definition of the uniform  $(a, b)$  PDF,

$$\text{Var}[X|Y = 1/2] = \frac{(1 - 1/2)^2}{12} = \frac{1}{48}.$$

## Section 7.6

---

Bivariate Gaussian Random  
Variables: Conditional PDFs

## Theorem 7.15

---

If  $X$  and  $Y$  are the bivariate Gaussian random variables in Definition 5.10, the conditional PDF of  $Y$  given  $X$  is

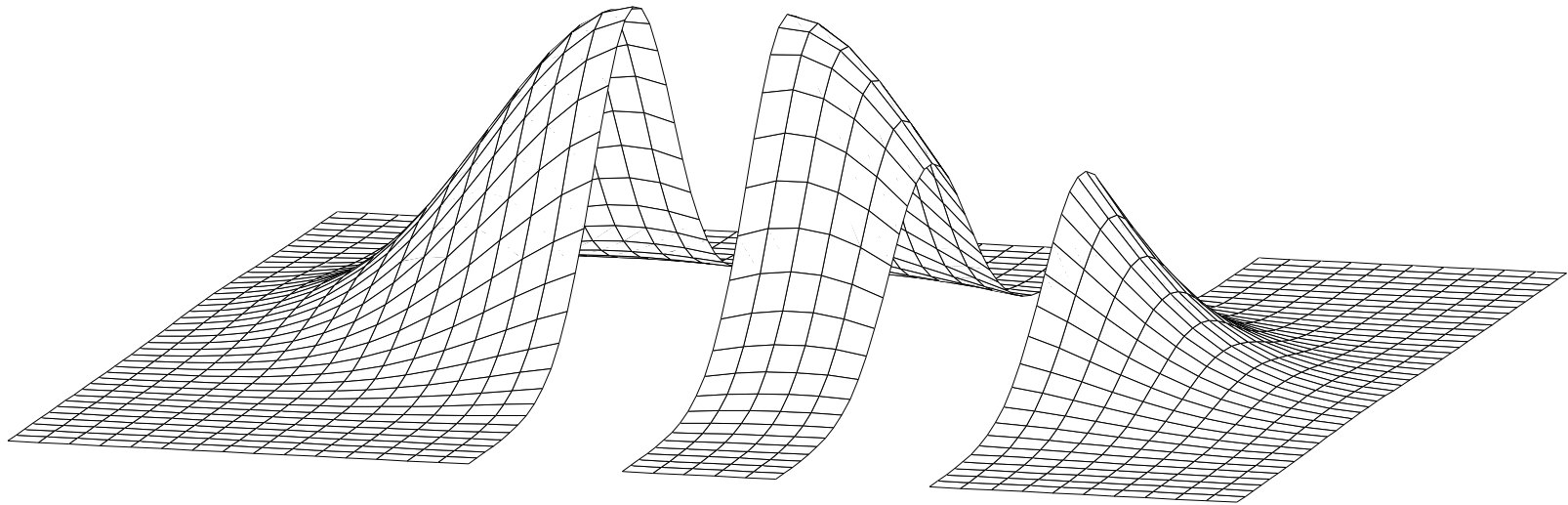
$$f_{Y|X}(y|x) = \frac{1}{\tilde{\sigma}_Y \sqrt{2\pi}} e^{-(y - \tilde{\mu}_Y(x))^2 / 2\tilde{\sigma}_Y^2},$$

where, given  $X = x$ , the conditional expected value and variance of  $Y$  are

$$\begin{aligned} \mathbb{E}[Y|X = x] &= \tilde{\mu}_Y(x) = \mu_Y + \rho_{X,Y} \frac{\sigma_Y}{\sigma_X} (x - \mu_X), \\ \text{Var}[Y|X = x] &= \tilde{\sigma}_Y^2 = \sigma_Y^2 (1 - \rho_{X,Y}^2). \end{aligned}$$

# Figure 7.1

---



Cross-sectional view of the joint Gaussian PDF with  $\mu_X = \mu_Y = 0$ ,  $\sigma_X = \sigma_Y = 1$ , and  $\rho_{X,Y} = 0.9$ . Theorem 7.15 confirms that the bell shape of the cross section occurs because the conditional PDF  $f_{Y|X}(y|x)$  is Gaussian.



## Theorem 7.16

---

If  $X$  and  $Y$  are the bivariate Gaussian random variables in Definition 5.10, the conditional PDF of  $X$  given  $Y$  is

$$f_{X|Y}(x|y) = \frac{1}{\tilde{\sigma}_X \sqrt{2\pi}} e^{-(x - \tilde{\mu}_X(y))^2 / 2\tilde{\sigma}_X^2},$$

where, given  $Y = y$ , the conditional expected value and variance of  $X$  are

$$\begin{aligned} \mathbb{E}[X|Y = y] &= \tilde{\mu}_X(y) = \mu_X + \rho_{X,Y} \frac{\sigma_X}{\sigma_Y} (y - \mu_Y), \\ \text{Var}[X|Y = y] &= \tilde{\sigma}_X^2 = \sigma_X^2 (1 - \rho^2). \end{aligned}$$

# Proof: Theorem 5.19

---

**(Theorem 5.19)** We define  $g(X, Y) = (X - \mu_X)(Y - \mu_Y)/(\sigma_X\sigma_Y)$ . From Definition 5.5 and Definition 5.6, we have the following formula for the correlation coefficient of any pair of random variables  $X$  and  $Y$ :

$$\mathbb{E}[g(X, Y)] = \frac{\mathbb{E}[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X\sigma_Y}. \quad (7.57)$$

We will now show that  $\mathbb{E}[g(X, Y)] = \rho_{X,Y}$  for bivariate Gaussian random variables  $X$  and  $Y$ . Using the substitution  $f_{X,Y}(x, y) = f_{Y|X}(y|x)f_X(x)$  to evaluate the double integral in the numerator, we obtain

$$\begin{aligned} \mathbb{E}[g(X, Y)] &= \frac{1}{\sigma_X\sigma_Y} \int_{-\infty}^{\infty} (x - \mu_X) \left( \int_{-\infty}^{\infty} (y - \mu_Y) f_{Y|X}(y|x) dy \right) f_X(x) dx \\ &= \frac{1}{\sigma_X\sigma_Y} \int_{-\infty}^{\infty} (x - \mu_X) \mathbb{E}[Y - \mu_Y|X = x] f_X(x) dx. \end{aligned} \quad (7.58)$$

Because  $\mathbb{E}[Y|X = x] = \tilde{\mu}_Y(x)$  in Theorem 7.15, it follows that

$$\mathbb{E}[Y - \mu_Y|X = x] = \tilde{\mu}_Y(x) - \mu_Y = \rho_{X,Y} \frac{\sigma_Y}{\sigma_X} (x - \mu_X). \quad (7.59)$$

Applying Equation (7.59) to Equation (7.58), we obtain

$$\mathbb{E}[g(X, Y)] = \frac{\rho_{X,Y}}{\sigma_X^2} \int_{-\infty}^{\infty} (x - \mu_X)^2 f_X(x) dx = \rho_{X,Y}, \quad (7.60)$$

because the integral in the final expression is  $\text{Var}[X] = \sigma_X^2$ .

# Gaussian Conditional Variances

- Theorem 5.14 states that for any pair of random variables,  $|\rho_{X,Y}| < 1$ .
- Introducing this inequality to the formulas for conditional variance in Theorem 7.15 and Theorem 7.16 leads to the following inequalities:

$$\text{Var} [Y|X = x] = \sigma_Y^2(1 - \rho_{X,Y}^2) \leq \sigma_Y^2, \quad (7.61)$$

$$\text{Var} [X|Y = y] = \sigma_X^2(1 - \rho_{X,Y}^2) \leq \sigma_X^2. \quad (7.62)$$

These formulas state that for  $\rho_{X,Y} \neq 0$ , learning the value of one of the random variables leads to a model of the other random variable with reduced variance.

## Quiz 7.6

---

Let  $X$  and  $Y$  be jointly Gaussian  $(0, 1)$  random variables with correlation coefficient  $1/2$ . What is the conditional PDF of  $X$  given  $Y = 2$ ? What are the conditional expected value and conditional variance  $E[X|Y = 2]$  and  $\text{Var}[X|Y = 2]$ ?

## Quiz 7.6 Solution

---

Since  $X$  and  $Y$  are bivariate Gaussian random variables with  $\rho = 1/2$ ,  $\mu_X = \mu_Y = 0$ , and  $\sigma_X = \sigma_Y = 1$ , Theorem 7.16 tells us that given  $Y = y$ ,  $X$  is conditionally Gaussian with parameters

$$\tilde{\mu}_X(y) = \rho y = \frac{y}{2}, \quad \tilde{\sigma}_X^2 = 1 - \rho^2. \quad (1)$$

For  $y = 2$ , we have

$$\tilde{\mu}_X = \tilde{\mu}_X(2) = 1 \quad \tilde{\sigma}_X^2 = 3/4. \quad (2)$$

The conditional PDF of  $X$  is

$$f_{X|Y}(x|2) = \frac{1}{\sqrt{2\pi\tilde{\sigma}_X^2}} e^{-(x-\tilde{\mu}_X)^2/2\tilde{\sigma}_X^2} = \frac{1}{\sqrt{3\pi/2}} e^{-2(x-1)^2/3}. \quad (3)$$

## Section 7.7

---

Matlab

## Example 7.17 Problem

---

Repeating Example 7.2, find the conditional PMF for the length  $X$  of a video given event  $L$  that the video is long with  $X \geq 5$  minutes.

## Example 7.17 Solution

---

```
sx=(1:8)';  
px=[0.15*ones(4,1);...  
    0.1*ones(4,1)];  
sxL=unique(find(sx>=5));  
pL=sum(finitepmf(sx,px,sxL));  
pxL=finitepmf(sx,px,sxL)/pL;
```

$x_i \in L$ .

With random variable  $X$  defined by  $sx$  and  $px$  as in Example 3.36, this code solves this problem. The vector  $sxL$  identifies the event  $L$ ,  $pL$  is the probability  $P[L]$ , and  $pxL$  is the vector of probabilities  $P_{X|L}(x_i)$  for each



## Example 7.18 Problem

---

Write a function `xy = xytrianglerv(m)` that generates  $m$  sample pairs  $(X, Y)$  in Example 7.13.

## Example 7.18 Solution

---

In Example 7.13, we found that

$$f_X(x) = \begin{cases} 2x & 0 \leq x \leq 1, \\ 0 & \text{otherwise,} \end{cases} \quad f_{Y|X}(y|x) = \begin{cases} 1/x & 0 \leq y \leq x, \\ 0 & \text{otherwise.} \end{cases} \quad (7.63)$$

```
function xy = xytrianglererv(m);  
x=sqrt(rand(m,1));  
y=x.*rand(m,1);  
xy=[x y];
```

For  $0 \leq x \leq 1$ , we have that  $F_X(x) = x^2$ . Using Theorem 6.5 to generate sample values of  $X$ , we define  $u = F_X(x) = x^2$ . Then, for  $0 < u < 1$ ,  $x = \sqrt{u}$ . By Theorem 6.5, if  $U$  is uniform  $(0, 1)$ , then  $\sqrt{U}$  has PDF  $f_X(x)$ . Next, we observe that given  $X = x_i$ ,  $Y$  is the uniform  $(0, x_i)$  random variable. Given another uniform  $(0, 1)$  random variable  $U_i$ , Theorem 6.3(a) states that  $Y_i = x_i U_i$  is the uniform  $(0, x_i)$  random variable. We implement these ideas in the function `xytrianglererv.m`.

## Quiz 7.7

---

For random variables  $X$  and  $Y$  with joint PMF  $P_{X,Y}(x,y)$  given in Example 7.9, write a Matlab function `xy=dtrianglererv(m)` that generates  $m$  sample pairs.

## Quiz 7.7 Solution

---

One straightforward method is to follow the approach of Example 5.26. Instead, we use an alternate approach. First we observe that  $X$  has the discrete uniform  $(1, 4)$  PMF. Also, given  $X = x$ ,  $Y$  has a discrete uniform  $(1, x)$  PMF. That is,

$$P_X(x) = \begin{cases} 1/4 & x = 1, 2, 3, 4, \\ 0 & \text{otherwise,} \end{cases} \quad P_{Y|X}(y|x) = \begin{cases} 1/x & y = 1, \dots, x, \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

Given  $X = x$ , and an independent uniform  $(0, 1)$  random variable  $U$ , we can generate a sample value of  $Y$  with a discrete uniform  $(1, x)$  PMF via  $Y = \lceil xU \rceil$ . This observation prompts the following program:

```
function xy=dtrianglerv(m)
sx=[1;2;3;4];
px=0.25*ones(4,1);
x=finiterv(sx,px,m);
y=ceil(x.*rand(m,1));
xy=[x';y'];
```